

Oral Probability Questions

These are notes made in preparation for oral exams involving the following topics in probability: Random walks, Martingales, and Markov Chains. Textbook used: "Probability: Theory and Examples," Durrett.

Chapter 4

1. Define: Random Walk

Let X_1, X_2, \dots be iid taking values in \mathbb{R}^d

and let $S_n = X_1 + \dots + X_n$. S_n is a random walk.

2. Name a Random Walk Theorem

- **RW Possibilities on \mathbb{R} :** Four possibilities, one w/prob = 1.
 - $S_n = 0 \forall n$, (recurrent)
 - $S_n \rightarrow \pm\infty$, (transient)
 - $-\infty = \liminf S_n < \limsup S_n = \infty$ (recurrent)
- **RW Recurrence on \mathbb{R}^d :**
 - S_n recurrent in $d=1$ if $S_n/n \rightarrow 0$ in probability. (or SSRW)
 - S_n recurrent in $d=2$ if S_n/\sqrt{n} converges in distribution to a non-deg. norm. dist. (or SSRW)
 - S_n transient in $d \geq 3$ if is "truly three-dimensional"
- **RW Equivalencies Theorem:** Let $\tau_0=0$ and $\tau_n = \inf\{m > \tau_{n-1} : S_m = 0\}$ be time of n th return to 0. Then, $P(\tau_1 < \infty) = 1 \Leftrightarrow P(S_m = 0 \text{ i.o.}) = 1 \Leftrightarrow \sum_{m=0}^{\infty} P(S_m = 0) = \infty$.
- **RW Convergence/Divergence Theorem:** Convergence (divergence) of $\sum_n P(|S_n| < \varepsilon)$, $\forall \varepsilon > 0$ is sufficient to determine transience (recurrence) of S_n .

3. Does (a version of 1) always have _____ property (related to 2)?

- For iid X_1, X_2, \dots , is exchangeable sigma field ε trivial? Yes. By Hewitt Savage 0-1. $P(A) \in \{0, 1\}$ for each $A \in \varepsilon$
- **Types of sets for RW recurrent values (V)?** Empty set, or a closed subgroup of \mathbb{R}^d .
- **If V (recurrent values) is a closed subgroup, $V = ?$** $V = \{\text{Possible Values}\}$

4. Question that leads to a Counterexample/Example.

- Are SSRW always recurrent? They are on $d < 3$.
- Are RW on \mathbb{R}^d always recurrent w/ $d < 3$? No, only w/ SSRW or w/ correct convergence (see above)
- **Will Wald's theorem hold with a SSRW $S_n = X_1 + \dots + X_n$, with $X_n \in \{\pm 1\}$ starting at $S_0 = 0$, with a stopping time T when $S_T = s \neq 0$?** (Wald has X_i as iid w/ $E[t] < \infty$ and $E[X_i] < \infty$)

Note that for any SSRW, that the time T to any position $S_T = s$ is finite, with probability one.

However, the expected time is infinite. Therefore, it does not satisfy one of Wald's Theorem's assumptions.

Proof by Contradiction: Having conditioned on $C = \{S_T = X_1 + \dots + X_T = s\}$, then the conditioned expectation $E(X_1 + \dots + X_T | C) = s$ is evident; furthermore, since $X_n = \pm 1$ for all n with equal probability, we easily see that $\mu = E(X_n) = 0$. Under these observations, assuming Wald's Identity ($E[S_T] = \mu \cdot T$), we obtain an immediate contradiction ($s \neq 0 \cdot T$).

- **If S, T are stopping times, then is it necessary that (S - T) is a stopping time?**

S-T is not necessarily a stopping time. For a counterexample, consider the simple random walk (X_n) on $\{\dots, -1, 0, 1, \dots\}$ starting at $X_0=0$, and let $S:=\inf\{n: X_n=1\}$ and $T:=1$. Note that $\{S-T=1\}=\{S=2\}$ which is not X_1 -measurable.

Examples of stopping times

- To illustrate some examples of random times that are stopping rules and some that are not, consider a gambler playing roulette with a typical house edge, starting with \$100 and betting \$1 on red in each game:
- Playing exactly five games corresponds to the stopping time $\tau = 5$, and is a stopping rule.
- Playing until he either runs out of money or has played 500 games is a stopping rule.
- Playing until he is the maximum amount ahead he will ever be is not a stopping rule and does not provide a stopping time, as it requires information about the future as well as the present and past.
- Playing until he doubles his money (borrowing if necessary) is not a stopping rule, as there is a positive probability that he will never double his money.
- Playing until he either doubles his money or runs out of money is a stopping rule, even though there is potentially no limit to the number of games he plays, since the probability that he stops in a finite time is 1.

1. Define: Stopping Time

$(\Omega, \mathcal{F}, (\mathcal{F}_n)_{n \geq 0}, \mathbb{P})$ a filtered prob space.

Stopping time $T : \Omega \rightarrow \mathbb{Z}_+ \cup \{+\infty\}$ is r.v. s.t. $\{T \leq n\} \in \mathcal{F}_n$

$\forall n \geq 0$, or equivalently, $\{T = n\} \in \mathcal{F}_n$ for all $n \geq 0$.

2. Name a Stopping Time Theorem

- **Wald's Identity:** Let X_1, X_2, \dots be iid w/ $\mu:=E[X_n]<\infty$. Set X_0 and let $S_n=X_1+\dots+X_n$, and T be stopping time w/ $E[T]<\infty$. Then, $E[S_T]=\mu E[T]$.
- If S, T, T_n are stopping times on $(\Omega, \mathcal{F}, \mathcal{F}_n, P)$. Then so are:
 - $S+T, S \wedge T:=\min(S, T), S \vee T:=\max(S, T)$
 - $\liminf_n T_n$ and $\inf_n T_n, \limsup_n T_n$ and $\sup_n T_n$

3. Does (a version of 1) always have _____ property (related to 2)?

- Are constants stopping times? Yes.

4. Question that leads to a Counterexample/Example.

If stopping time T and \mathcal{F}_T , and X_1, X_2, \dots iid, is $\{X_{T+n}\}_{n>0}$ independent of \mathcal{F}_T for all T? Yes.

Examples of Stopping Times:

- Constants
- If X_n is an adapted process, and $A \in \mathcal{F}$, The first entry time into A is a stopping time.

Chapter 5

1. Define: Martingale (or sub, or super)

X_n on $(\Omega, \mathcal{F}, P, \mathcal{F}_n)$, s.t.

- X_n is adapted to \mathcal{F}_n .
- $E|X_n| < \infty$ for each n .
- $E[X_{n+1} | \mathcal{F}_n] = X_n$ a.s. $\forall n$. (or \geq , or \leq resp.)

2. Name a Martingale Theorem

- **Stopping Time (Super)Martingale Prop:** If T is a stopping time and X_n is a (super)mart, then $X_{T \wedge n}$ is a (super)mart.
- **Submartingale Convergence:** Suppose that X_n is a sub-martingale with $\sup_n E[X_n^+] < \infty$. Then for some X , we have $X_n \rightarrow X$ a.s., where $E|X| < \infty$.
- **Martingale Convergence:** If X_n is a martingale with $\sup_n E|X_n| < \infty$, then $X_n \rightarrow X$ a.s. and $E|X| < \infty$.
- **Nonnegative SuperMartingale Convergence:** If X_n is a super-martingale with $X_n \geq 0$, then $X_n \rightarrow X$ a.s. and $E[X] \leq E[X_0]$
- **Galton-Watson:** Let $\xi_i^n, i \geq 1, n \geq 0$ be iid nonnegative integer-valued r.v.s with a common $\mu = E[\xi_i^n] \in (0, \infty)$. Define $Z_0 = 1$ and $Z_{n+1} = \{\xi_1^{n+1} + \dots + \xi_{Z_n}^{n+1}$ if $Z_n > 0$; and 0 if $Z_n = 0$. Then, $(Z_n / \mu^n)_{n \geq 0}$ is a martingale with respect to $\mathcal{F}_n = \sigma(\xi_i^m : i \geq 1, 0 \leq m < n)$.

3. Does (a version of 1) always have _____ property (related to 2)?

- *Do supermartingales always converge a.s.?* Not necessarily, it's guaranteed when X_n nonnegative.
- *If $\mu < 1$, Then $P(\text{extinction}) = ?$* $P(\text{extinction}) = 1$.

4. Question that leads to a Counterexample/Example.

- **When $\mu = 1$, is $P(\text{extinction})$ equal to 1?** Only when $P(\xi_1 = 1) < 1$.
- **From Durrett Exmpl. 5.2.3: Do nonnegative martingales converge in L^1 ?**
Not always. Let S_n be a symmetric simple random walk with $S_0 = 1$, i.e., $S_n = S_{n-1} + \xi_n$ where ξ_1, ξ_2, \dots are i.i.d. with $P(\xi_i = 1) = P(\xi_i = -1) = 1/2$. Let $N = \inf\{n : S_n = 0\}$ and let $X_n = S_{N \wedge n}$. Since the martingale property is closed under stopping times, X_n is a nonnegative martingale. The Nonnegative SuperMartingale Convergence Theorem implies X_n converges a.s. to a limit $X_\infty < \infty$ that must be $\equiv 0$, since convergence to $k > 0$ is impossible. (If $X_n = k > 0$ then $X_{n+1} = k \pm 1$.) Since $E X_n = E X_0 = 1$ for all n and $X_\infty = 0$, convergence cannot occur in L^1 . $E|X_n - X_\infty| = E[X_n] \rightarrow 1 \neq 0$.
- **Consider the random walk $S_n = X_1 + \dots + X_n$ starting at zero with X 's having $P(X_i = 1) = P(X_i = -1) = 1/2$, a martingale. Now if $T = \inf\{n \geq 0 : S_n = 1\}$. Can we bound T ?**
No. For any $n \in \{1, 2, \dots\}$ we have $P(S_k \leq 0 \text{ for all } k \leq n) \geq P(X_1 = \dots = X_n = -1) = 1/2^n$ since $\{S_k \leq 0 \text{ for all } k \leq n\} \subseteq \{T > n\}$, this implies $P(T > n) \geq P(S_k \leq 0 \text{ for all } k \leq n) \geq 1/2^n > 0$. As $n \in \mathbb{N}$ is arbitrary, this proves that T is unbounded.
- **Do all Martingales which converge in probability, also do so in L^1 ?**
No. Any martingale which converges almost surely but not in L^1 does the job (since a.s. conv. implies conv. in prob.); see example 5.2.3 above.

- If $E(X_{n+1}|X_n)=X_n$ for all n , must X_n be a martingale (instead of $E(X_{n+1}|F_n)=X_n$)?

No. Let $(Y_j)_{j \in \mathbb{N}}$ be a sequence of iid r.v. such that $EY_j=0$. Fix $N \in \{1,2,\dots\}$ and define: $X_n := \sum_{j=1}^n Y_j$ for all $n \leq N$, and $X_n := \sum_{j=1}^N Y_j + Y_1 - Y_2 = X_N + Y_1 - Y_2$ for all $n > N$.

For $n \leq N$ and $n > N+1$, the condition $E(X_n|X_{n-1})=X_{n-1}$ is obviously satisfied.

For $n=N+1$, we have $E(X_{N+1}|X_N)=X_N+E(Y_1|X_N)-E(Y_2|X_N)$. Since $(Y_j)_{j \in \mathbb{N}}$ is identically distributed and independent, we have $E(Y_1|X_N)=E(Y_2|X_N)$ and therefore $E(X_{N+1}|X_N)=X_N$. On the other hand,

$$\begin{aligned} \mathbb{E}(X_{N+1} | \mathcal{F}_N) &= X_N + \underbrace{2\mathbb{E}(Y_1 | \mathcal{F}_N)}_{\mathbb{E}(X_1|\mathcal{F}_N)=X_1} - \underbrace{\mathbb{E}(Y_1 + Y_2 | \mathcal{F}_N)}_{\mathbb{E}(X_2|\mathcal{F}_N)=X_2} = X_N + 2Y_1 - (Y_1+Y_2) \\ &= X_{N+1} \neq X_N. \end{aligned}$$

So, X_n is not a martingale.

1. Define: Optional Stopping Sigma-Field

Let $(\Omega, \mathcal{F}, (\mathcal{F}_n)_{n \geq 0}, \mathbb{P})$ and T be stopping time.

Denote by \mathcal{F}_T , the σ -field of "events which occur prior to time T ."

In symbols: $\mathcal{F}_T := \{A \in \mathcal{F} : A \cap \{T \leq n\} \in \mathcal{F}_n, \forall n \geq 0\}$.

2. Name an Optional Stopping Time Theorem

Optional Stopping Thm
for SubMarts

(or mart)

If S, T are stopping times w/ $\mathbb{P}(S \leq T < \infty) = 1$,

and $(X_{T \wedge n})_{n \geq 0}$ is UI submart, then $\mathbb{E}[X_T | \mathcal{F}_S] \geq X_S$ a.s.

Consequently, $\mathbb{E}[X_S] \leq \mathbb{E}[X_T]$. (switch to $=$'s for mart)

3. Does (a version of 1) always have _____ property (related to 2)?

- If T is a stopping time, then is \mathcal{F}_T a Sigma field? Yes
- If X_n is UI sub-martingale and T a stopping time, is $X_{T \wedge n}$ UI? Yes
- If $S \leq T$ are stopping times, then is $\mathcal{F}_T \subseteq \mathcal{F}_S$? No, but $\mathcal{F}_S \subseteq \mathcal{F}_T$.

4. Question that leads to a Counterexample/Example.

- If T is a stopping time, and X_n adapted, then is $X_T \in \mathcal{F}_T$? Not necessarily, this is only guaranteed when $P(T < \infty) = 1$.

1. Define: Conditional Expectation

(Ω, \mathcal{F}, P) w/ $X \in L^1$, $G \subseteq \mathcal{F}$, $Y := \mathbb{E}[X|G]$ is unique s.t.

Y is G -measurable and $\mathbb{E}|Y| < \infty$.

$$\mathbb{E}[\mathbb{E}[X|G]1_A] = \mathbb{E}[Y1_A] = \mathbb{E}[X1_A], A \in G$$

2. Name a Conditional Expectation Theorem

- **Conditional MCT:** Let $G \subseteq \mathcal{F}$.

Let $X, X_n \geq 0$ be integrable r.v.s and $X_n \uparrow X$.

Then $\mathbb{E}[X_n|G] \uparrow \mathbb{E}[X|G]$ a.s.

- **Conditional DCT:** Let $G \subseteq \mathcal{F}$.

If $X_n \rightarrow X$ a.s. and $|X_n| \leq Y$ for some integrable r.v. Y .

Then $\mathbb{E}[X_n|G] \rightarrow \mathbb{E}[X|G]$ a.s.

- **Conditional Jensen's:** Let $G \subseteq \mathcal{F}$.

If $\varphi : \mathbb{R} \rightarrow \mathbb{R}$ is convex, $\mathbb{E}|X| < \infty$ and $\mathbb{E}|\varphi(X)| < \infty$,

then $\mathbb{E}[\varphi(X)|G] \geq \varphi(\mathbb{E}[X|G])$ a.s.

3. Does (a version of 1) always have _____ property (related to 2)

4. Question that leads to a Counterexample/Example.

- If X, Y are two random variables and $E(X|Y)=E(X)$, are X and Y independent?

Not necessarily. Let $X \in \{-1, 0, 1\}$, each with probability $\frac{1}{3}$. Let $Y=X^2$. Note that X and Y are not independent. However, observe that $E(X|Y=0)=0$ and $E(X|Y=1) = \frac{1}{3}(-1) + \frac{1}{3}(1) = 0$, so $E(X|Y)=0=E(X)$ with probability 1.

1. Define: Uniform Integrability

Family of r.v.s $(X_\alpha)_{\alpha \in \Lambda}$ is uniformly integrable (UI) if

$$\sup_{\alpha \in \Lambda} \mathbb{E}[|X_\alpha| 1_{\{|X_\alpha| > M\}}] \rightarrow 0 \text{ as } M \rightarrow \infty.$$

Remrk: Since $\mathbb{E}|X_\alpha| \leq M + \mathbb{E}[|X_\alpha| 1_{\{|X_\alpha| > M\}}]$, then $UI \Rightarrow L^1$ -bounded uniformly for $(X_\alpha)_{\alpha \in \Lambda}$.

2. Name a UI Theorem

- **Sub σ -field UI Lemma:** Let $X \in L^1(\Omega, \mathcal{F}, P)$. Then, $\{\mathbb{E}[X|G] : G \text{ a } \sigma\text{-field } \subset \mathcal{F}\}$ is UI. Used in Levy's Fwd Law.
- **If $X_n \rightarrow X$ in probability, then TFAE:**
 - $\{X_n\}$ is UI.
 - $X_n \rightarrow X$ in L^1 . $\mathbb{E}|X_n - X| \rightarrow 0$.
 - $\mathbb{E}|X_n| \rightarrow \mathbb{E}|X| < \infty$.
- **Convergence in Prob Corollary:**
 - If $X_n \rightarrow X$ in prob. and $\{X_n\}$ is UI $\Leftrightarrow X_n \rightarrow X$ in L^1 .
 - If $X_n \rightarrow X$ in prob and $|X_n| \leq Y$ for some $Y \in L^1$ (L^1 bounded), then $X_n \rightarrow X$ in L^1 .
- **Submartingale Equivalencies Thm:** For a submart X_n , TFAE:
 - $\{X_n\}$ is UI.
 - X_n converges a.s. and in L^1 .
 - X_n converges in L^1 .
 - If X_n is a martingale, then \exists integrable r.v. X so that $X_n = \mathbb{E}[X|F_n]$.

3. Does (a version of 1) always have _____ property (related to 2)?

- Do UI sub martingales converge almost surely? Yes.

4. Question that leads to a Counterexample/Example.

- For a reverse martingale $(X_n)_n$, clearly, $E[X_0]=X_n$, for each $n \in \{1,2,\dots\}$. Is $E[X_0 | F_n]$ UI?

Yes. Proof: Since $(X_n)_n$ is a martingale, we have: $E[X_0] < \infty$. So by the Subsigma Field UI Lemma, we have $E[X_0 | F_n]$ is UI.

- Durrett Example 5.5.1. Suppose X_1, X_2, \dots are UI and $X_n \rightarrow X$ a.s. Need $E(X_n|F)$ converge a.s.?

No. Let Y_1, Y_2, \dots and Z_1, Z_2, \dots be independent r.v.'s with $P(Y_n = 1) = 1/n$, $P(Y_n = 0) = 1 - 1/n$, $P(Z_n = n) = 1/n$, $P(Z_n = 0) = 1 - 1/n$. So our counterexample uses $X_n := Y_n Z_n$. Observe that $E(X_n : |X_n| \geq 1) = n/n^2$, so X_n is UI. Also, $P(X_n > 0) = 1/n^2$ so $\sum P(X_n > 0) < \infty$, $P(\{X_n > 0\} \text{ i.o.}) = 0$, and the Borel-Cantelli lemma implies $X_n \rightarrow 0$ a.s. Let $F = \sigma(Y_1, Y_2, \dots)$. Then, $E(X_n|F) = Y_n E(Z_n|F) = Y_n E(Z_n) = Y_n$. Since $Y_n \rightarrow 0$ in L^1 but not a.s., the same is true for $E(X_n|F)$. Since $\sum P\{Y_n > 1/2\} = \sum 1/n = \infty$. Then, apply Borel-Cantelli.

- Does every sequence X_n which converges almost surely, also converge in L^1 ?

No, take the sequence $n \cdot 1_{[0, 1/n]}$, and note that it converges almost surely to zero. Also note that $E[n \cdot 1_{[0, 1/n]}] = 1$ for all n . So, $\lim E[n \cdot 1_{[0, 1/n]} - X] = \lim E[n \cdot 1_{[0, 1/n]}] = 1 \neq 0$.

- For a martingale X_n does UI imply integrability of $\sup|X_n|$?

No, but the counterexamples are not trivial.

- Non-trivial martingale which converges almost surely to 0

Let Y_1, Y_2, \dots be nonnegative i.i.d. random variables with $E[Y_m] = 1$ and $P(Y_m = 1) < 1$.

(i) Show that $X_n = \prod_{m \leq n} Y_m$ defines a martingale. (ii) Use an argument by contradiction to show $X_n \rightarrow 0$ a.s.

(i) is easy to check.

(ii) Let $X = \lim X_n$. The Hewitt-Savage zero one law says (since $X \in \{\text{exchangeable sigma field}\}$) that X is almost surely a constant. Also, $X = Y_1 \cdot \prod_{i=2}^{\infty} Y_i$ has the same distribution as $Y_1 \cdot X$. Since Y_1 is not constant a.s., this forces $X \in \{0, \infty\}$, but $X \neq \infty$ since by Fatou and Y_n independence we have: $E(X) = E(\lim X_n) = E(\lim \prod_{m \leq n} Y_m) \leq \lim E(\prod_{m \leq n} Y_m) = \lim \prod_{m \leq n} E(Y_m) = 1$. So $X = 0$, and $X_n \rightarrow 0$ a.s.

Chapter 6

1. Define: Markov Chain

An $\{F_n\}$ -adapted stochastic process X_n taking values in (S, \mathcal{S}) is called a Markov chain if it has the **Markov Property**: $P(X_{n+1} \in B | F_n) = P(X_{n+1} \in B | X_n)$ a.s. for each $B \in \mathcal{S}$, $n \geq 0$.

2. Name a Markov Chain Theorem

- **Decomposition Theorem**: Let $R = \{x : \rho_{xx} = 1\}$ be the recurrent states of a Markov chain. R can be written as $\cup_i R_i$, where each R_i is closed and irreducible. [This results shows that for the study of recurrent states we can, without loss of generality, consider a single irreducible closed set.]
- **For an irreducible and recurrent chain (Corolary 6.46)**:
 - The stat/inv measures are unique up to constant multiples.
 - If μ is a stat/inv measure, then $\mu(x) > 0$ for all x .
- **If p is irreducible and has a stationary distribution π** .
 - **Calculating Stat/Inv Distribution**: $\pi(x) = 1/E_x[T_x]$.
 - **Theorem D6.5.7**: Any other stationary measure is a multiple of π .

- **Theorem 6.70** (Markov Chain Convergence Theorem): Consider an irreducible, aperiodic Markov chain with stationary distribution π . Then, $p^n(x,y) \rightarrow \pi(y)$ as $n \rightarrow \infty$, for all $x,y \in S$.
- **Theorem 6.62** (Asymptotic Density of Returns): Let $y \in S$ be recurrent, and $N_n(y) = \sum_{i=1}^n 1_{\{X_i=y\}}$, then $\lim N_n(y)/n = (1/E_y[T_y]) 1_{\{T_y < \infty\}}$, P_x - a.s.

3. Does (a version of 1) always have _____ property (related to 2)?

4. Question that leads to a Counterexample/Example.

- **Multivalued Markov Chain:** If ξ_0, ξ_1, \dots are iid $\in \{H, T\}$, each with $p=1/2$, then $X_n := \{\xi_n, \xi_{n+1}\}$ is a Markov chain.
- **(HW 3):** If ξ_0, ξ_1, \dots are iid $\in \{-1, 1\}$ with $p=1/2$, and $S_0=0$, $S_n := \xi_1 + \xi_2 + \dots + \xi_n$, and $X_n = \max\{S_m : 0 \leq m \leq n\}$. Then is X_n a Markov chain?

No. Observe the sequence $(X_1, X_2, X_3) = (1, 1, 1)$. This can occur if $(S_1, S_2, S_3) = (1, 0, 1)$, or if $(S_1, S_2, S_3) = (1, 0, -1)$. Therefore, we have: $P(X_4=2 | X_1=1, X_2=1, X_3=1) = (1/2) \cdot (1/2) = 1/4$. Alternatively, take the sequence $(X_1, X_2, X_3) = (0, 0, 1)$, and observe that this only occurs in only one way, namely if $(S_1, S_2, S_3) = (-1, 0, 1)$. Therefore, $P(X_4=2 | X_1=0, X_2=0, X_3=1) = 1 \cdot (1/2) = 1/2$. Therefore, since the dependence includes more than just the previous value, X_n is not a Markov chain.

1. **Define: Stationary Distribution**

It's a stationary/invariant measure that is also a probability measure: $\pi p = \pi$ such that $\pi(y) = \sum_{x \in S} \pi(x) p(x,y)$, and $\sum_{x \in S} \pi(x) = 1$. It represents a possible equilibrium for the chain.

2. **Name a Stationary Distribution Theorem**

- **If p is irreducible and has a stationary distribution π .**
 - **Calculating Stat/Inv Distribution:** $\pi(x) = 1/E_x[T_x]$.
 - **Theorem D6.5.7:** Any other stationary measure is a multiple of π .
- **Recurrence from Positive Stat/Inv Distributions:** If π is a stationary/invariant distribution of a Markov chain and $\pi(x) > 0$ for some x , then that x is recurrent.
- **Theorem 6.70** (Markov Chain Convergence Theorem): Consider an irreducible, aperiodic Markov chain with stationary distribution π . Then, $p^n(x,y) \rightarrow \pi(y)$ as $n \rightarrow \infty$, for all $x,y \in S$.

3. Does (a version of 1) always have _____ property (related to 2)?

- **What are sufficient conditions for a Markov chain's stat/inv measures to be unique up to constant multiples?** That it be irreducible and recurrent.
- **What are sufficient conditions for a Markov chain's stat/inv measure, if it exists, to have the property $\mu(x) > 0$ for all x ?** That it be irreducible and recurrent.
- **What are sufficient conditions for a Markov chain's stat/inv distribution, if it exists, to be unique?** That it be irreducible and recurrent.
- **Assume a Markov chain is irreducible and recurrent, what are sufficient conditions to allow us to conclude that the stat/inv distribution cannot exist?** The stat/inv measure has infinite mass.
- **If π is a stat/inv distribution and $\pi(x) > 0$, what we know about x ?** It is recurrent.
- **If you have an irreducible Markov chain, and there is a positive recurrent value, does this imply the existence of a stationary distribution?** Yes.
- **If you have an irreducible Markov chain, and every state is positive recurrent, does this imply the existence of a stationary distribution?** Yes.
- **If you have an irreducible Markov chain that has a stationary distribution, does this imply the existence of a positive recurrent value?** Yes.

4. **Question that leads to a Counterexample/Example.**

- Let X_n be a Markov chain, where S is the state space and P is the transition matrix. Is every closed class recurrent? No, for example a biased random walk on the integers is transient. *Finite* closed classes, on the other hand, are always recurrent.

1. **Define: Markov Chain Recurrence**

A state $y \in S$ is called recurrent if $\rho_{yy} = 1$, and is called transient if $\rho_{yy} < 1$.

2. **Name a Recurrence Theorem**

- **Decomposition Theorem:** Let $R = \{x : \rho_{xx} = 1\}$ be the recurrent states of a Markov chain. R can be written as $\cup_i R_i$, where each R_i is closed and irreducible. [This results shows that for the study of recurrent states we can, without loss of generality, consider a single irreducible closed set.]
- **Theorem 6.62 (Asymptotic Density of Returns):** Let $y \in S$ be recurrent. Then $\lim N_n(y)/n = (1/E_y[T_y])1_{\{T_y < \infty\}}$, P_x - a.s.

3. **Does (a version of 1) always have _____ property (related to 2)?**

4. **Question that leads to an Counterexample/Example.**

1. **Define: Markov Chain Irreducibility**

Markov chain is irreducible if it is possible to get to any state from any state. Formally, if its state space is a single communicating class, i.e., $x \leftrightarrow y$ for all $x, y \in S$.

2. **Name an Irreducibility Theorem**

- **Decomposition Theorem:** Let $R = \{x : \rho_{xx} = 1\}$ be the recurrent states of a Markov chain. R can be written as $\cup_i R_i$, where each R_i is closed and irreducible. [This results shows that for the study of recurrent states we can, without loss of generality, consider a single irreducible closed set.]
- **For an irreducible and recurrent chain (Corollary 6.46):**
 - The stat/inv measures are unique up to constant multiples.
 - If μ is a stat/inv measure, then $\mu(x) > 0$ for all x .
- **If p is irreducible and has a stationary distribution π .**
 - **Calculating Stat/Inv Distribution:** $\pi(x) = 1/E_x[T_x]$.
 - **Theorem D6.5.7:** Any other stationary measure is a multiple of π .
- **Theorem 6.70 (Markov Chain Convergence Theorem):** Consider an irreducible, aperiodic Markov chain with stationary distribution π . Then, $p^n(x, y) \rightarrow \pi(y)$ as $n \rightarrow \infty$, for all $x, y \in S$.

3. **Does (a version of 1) always have _____ property (related to 2)?**

4. **Question that leads to a Counterexample/Example.**

- If an irreducible Markov chain has period 2, then for every state $i \in S$ do we have $(P_{ii})^2 > 0$? No, consider $P =$

$$\begin{vmatrix} 0 & 1 \\ 1 & 0 \end{vmatrix}$$

Note that $P^2 = Id$, so period=2 and $x \leftrightarrow y$. So it is irreducible. But, $P_{ii} = 0$, so $(P_{ii})^2 = 0$.

Other Counterexamples/Examples

- **Are Martingales always Markov processes?**

No, assume that $(Z_t)_{t \geq 2}$ are independent, integrable, nonconstant (say, standard normal distributions), $\mu=0$, and Z_t independent of some X_0 , where $X_0:=X_1:=1$ and $X_t:=X_{t-1}+Z_t X_{t-2}$ for every $t \geq 2$. $F_n = \sigma\{X_1, \dots, X_n\}$. Then $E[X_t | F_{t-1}] = E[X_{t-1} | F_{t-1}] + E[Z_t X_{t-2} | F_{t-1}] = X_{t-1} + X_{t-2} E[Z_t | F_{t-1}] = X_{t-1}$ for every $t \geq 1$ (hence, if X_0 is integrable, $(X_t)_{t \geq 0}$ is a martingale) but $(X_t)_{t \geq 0}$ is not a Markov process since the conditional distribution of X_t on F_{t-1} does not depend on X_{t-1} only, but on (X_{t-1}, X_{t-2}) .

- **If X_n is a homogeneous Markov chain, is it true that X_{n^2} is also a homogeneous Markov chain?**

No. Consider the random walk on $\{\dots, -1, 0, 1, \dots\}$ that with probability $\frac{1}{3}$ each either: stays at its position, goes to the right, or to the left. We consider the particular transition probability:

$p^n(0,2) := P(X_{n^2} = 2 | X_{(n-1)^2} = 0)$, which if X_n is a homogeneous Markov chain, should not depend on n . But guess what? It depends on n . We have $p^1(0,2) = P(X_1 = 2 | X_0 = 0) = 0$, while $p^2(0,2) = P(X_4 = 2 | X_1 = 0) > 0$.

- **If $X_n \in \{-1, 1\}$, $S_0 = 0$, and $S_n := X_1 + \dots + X_n$. Then is $(|S_n|)_{n \geq 0}$ a Markov-chain?**

Not necessarily. Let $F_n = \sigma\{X_1, \dots, X_n\}$. It is not a Markov chain unless $p = \frac{1}{2}$ (probability of a step to the left), and a counterexample is to take $n=1$; then $|S_1|=1$ but $P(|S_2|=2) = p \neq \frac{1}{2}$ if the first step was to $S_1 = -1$, but is $P(|S_2|=2) = 1-p \neq \frac{1}{2}$ if the first step was to $S_1 = +1$. So, $P(|S_2|=2 : F_1) \in \{p, 1-p\}$ is not equal to $P(|S_2|=2 : |S_1|) = \frac{1}{2}(1-p) + \frac{1}{2}p = \frac{1}{2} \notin \{p, 1-p\}$, and $(|S_n|)_{n \geq 0}$ is not a Markov-chain

- **Does every chain that has a stationary distribution have a limiting distribution? No.**

Recall that a Markov chain has a limiting distribution if $\pi_j = \lim_{n \rightarrow \infty} p_{ij}^n, \forall i \in S$, exists. In particular, if the limit does not depend on the starting state (and hence distribution) of the chain.

We know a Markov Chain $\{X_n\}$ with a stat. distrib. μ as its initial distribution is a stationary process, because if $X_0 \sim \mu$ is a stationary distribution, then for each $n, X_n \sim \mu_{p_{n-1}} = \mu$. So, $(X_0, X_1, \dots, X_n) \sim (X_m, X_{m+1}, \dots, X_{m+n})$. Durrett said a special case to keep in mind for counterexamples is the Markov chain: $X_n: \Omega \rightarrow S = \{0, 1\}$ with transition probability $p(0,1) = p(1,0) = 1$, and stationary distribution $\mu(0) = \mu(1) = \frac{1}{2}$. Now let $X_0 \in \{0, 1\}$ w/probability $\frac{1}{2}$ (so not starting with the stat. dist.), so $(X_0, X_1, \dots) = (0, 1, 0, \dots)$ or $(1, 0, 1, \dots)$ with probability $\frac{1}{2}$. Note that it does not have a limiting distribution. Durrett is demonstrating that this chain satisfies stationarity, and that it is useful to keep this Markov chain in mind when *picturing* what stationarity means. In particular this is a commonly used counterexample to distinguish between stationary distributions, and limiting distributions.

Regarding the limiting distribution, note that in this case $\lim_{n \rightarrow \infty} p_{01}^n = 1$ and $\lim_{n \rightarrow \infty} p_{11}^n = 0$, so the limit does not exist. Any chain that has a limiting distribution necessarily is stationary (since π can be shown to satisfy the stationarity property). The converse however is not true: and this is what the counterexample shows, since the limit above only exists if the chain is started from $\mu(0) = \mu(1) = 1/2$, and not from an arbitrary distribution. In general for finite, irreducible Markov chains

- A stationary distribution always exists.
- Existence of a limiting distribution implies stationarity.
- If, in addition to being finite and irreducible, the chain is also aperiodic, then a limiting distribution is guaranteed to exist.